

Homework

Introduction to Econometrics

Fall 2005

Results

Using the 526 observations on workers available from Wooldrige's site in the Wage1 data set the following results are obtained. Consider the following linear model of log wages:

$$\ln(\text{wage}_i) = \beta_1 + \beta_2 \text{educ}_i + \beta_3 \text{exper}_i + \beta_4 \text{exper}^2 + \beta_5 \text{female}_i + u_i \quad (1)$$

where $i = 1, 2, \dots, 526$.

Least squares estimates

$$\begin{aligned} \widehat{\ln \text{wage}} &= 0.390483 + 0.0841361 \text{educ} + 0.0389100 \text{exper} - 0.000686022 \text{expersq} \\ &\quad \begin{matrix} (3.820) & (12.094) & (8.067) & (-6.389) \end{matrix} \\ &\quad - 0.337187 \text{female} \\ &\quad \quad \quad (-9.283) \\ T &= 526 \quad \bar{R}^2 = 0.3950 \quad F(4, 521) = 86.685 \quad R^2 = 0.3996 \\ &\quad \quad \quad (t\text{-statistics in parentheses}) \end{aligned}$$

1. R^2 is 0.3996.
2. Inclusion of exper and expersq allows the modeling of diminishing returns to experience. In this case, we expect $\beta_3 > 0$ and $\beta_4 < 0$.
3. $\partial E[\ln(\text{wage})]/\partial \text{educ} = \beta_2 = 0.0841$. Thus, an additional year of schooling results in an 8.41% increase in wages.
4. The RESET test statistic is: $F_{2,519} = 7.658522$, with p-value = $P(F_{2,519} > 7.65852) = 0.000527 < \alpha = 0.05$. The equation is misspecified at the 5% level of significance.
5. Adding educ^2 to the model and rerunning RESET yields: Test statistic: $F = 1.688653$, with p-value = $P(F_{2,518} > 1.68865) = 0.186$. The null hypothesis of no misspecification is not rejected at the 5% level. Things have certainly improve.
6. $\partial E[\ln(\text{wage})]/\partial \text{educ} = -0.0306529 + 2 * 0.00487164 * \text{educ}$. Therefore if Jethro has 6 years of schooling, the marginal effect is .0278. The marginal effect here is positive and so Jethro will benefit from finishing another year of schooling.
7. The estimated return to schooling for someone with a high school degree is marginal effect having 12 years of schooling is $-0.0306 + 2 * 0.00487 * 12 = 0.0862$. For a person with a college degree it is $-0.0306 + 2 * 0.00487 * 16 = 0.125$.

8. Test the null hypothesis that $\beta_5 = 0$ against the alternative $\beta_5 < 0$ using the t-ratio.

$$\widehat{\text{lwage}} = 1.01702 - 0.0306529 \text{educ} + 0.00487164 \text{educ}^2 + 0.0398883 \text{exper} \\ - 0.000719177 \text{expersq} - 0.319578 \text{female} \\ T = 526 \quad \bar{R}^2 = 0.4107 \quad F(5, 520) = 74.182 \quad \hat{\sigma} = 0.40803 \\ \text{(t-statistics in parentheses)}$$

The t-ratio is $-8.84 < -1.645$ and the hypothesis of no discrimination is rejected at the 5% level.

9. The augmented model is

$$\ln(\text{wage}_i) = \beta_1 + \beta_2 \text{educ}_i + \beta_3 \text{educ}_i^2 + \beta_4 \text{exper}_i \\ + \beta_5 \text{exper}_i^2 + \beta_6 \text{female}_i + u_i \quad (2)$$

where $i = 1, 2, \dots, 526$. The auxiliary model of the marginal effect of being female is

$$\beta_{6i} = \alpha_1 + \alpha_2 \text{married}_i \quad (3)$$

Substitution into the model (2) introduces an interaction term $\text{female}_i * \text{married}_i$ into the model. Its coefficient is

$$\ln(\text{wage}_i) = \beta_1 + \beta_2 \text{educ}_i + \beta_3 \text{educ}_i^2 + \beta_4 \text{exper}_i + \beta_5 \text{exper}_i^2 \\ + \alpha_1 \text{female}_i + \alpha_2 \text{female}_i * \text{married}_i + u_i \quad (4)$$

where $i = 1, 2, \dots, 526$. The coefficient on $\alpha_2 = -2.215$ which is significantly negative at the 5% level (it is less than -1.645).

10. Reformulating the model to allow discrimination for married women only (not single women) amounts to omitting the female_i in equation (4). This is probably **not** a good idea since α_1 was significant. Also, the coefficients change quite a bit when female is omitted—a likely sign of omitted variable bias.
11. The results table follows.

OLS estimates using the 526 observations. Dependent variable: lwage

Variable	Model (a)	Model (e)	Model (i)	Model (j)
const	0.390 (0.102)	1.017 (0.191)	0.968 (0.191)	0.8618 (0.195)
educ	0.0841 (0.00695)	-0.0306 (0.0305)	-0.0278 (0.0304)	-0.0367 (0.0312)
educ2	- -	0.00487 (0.00126)	0.00479 (0.00125)	0.0053 (0.00128)
exper	0.0389 (0.00482)	0.0399 (0.00476)	0.0426 (0.00491)	0.0480 (0.0049)
exper2	-0.000686 (0.000107)	-0.000719 (0.000106)	-0.00076 (0.000108)	-0.000859 (0.000109)
female	-0.337 (0.0363)	-0.319 (0.0361)	-0.255 (0.0461)	- -
marrfem	- -	- -	0.119 (0.0537)	-0.305 (0.0430)
RESET ¹	7.65	1.689	1.447	0.912

(asymptotic standard errors in in parentheses)

Model (i) is the preferred model. All coefficients except the one on *educ* are significantly different from zero. There is evidence from RESET that *educ*² should be included in the model. Also the married/female interaction term is significant and should not be omitted. The $F_{2,\infty}$ 5% critical value is 3.00.