

# Counts

- # Times an event occurs
- # patents
- visits

Could be zero for a substantial portion of popl.

Responses are discrete, ordered. But

- outcomes have cardinal value
- no upper bound.

## Poisson & Negative Binomial Models

$$y_i = \{0, 1, 2, \dots\}$$

$$x_i^T = \{1, x_{i2}, \dots, x_{ik}\} \text{ set of charact.}$$

$$\text{Assume } E(y_i | x_i) = \exp\{x_i^T \beta\} \quad (1)$$

This relates the AVG outcome to the characteristics. To be able to derive probabilities, we'll need a distrib.

## Several choices

- (1) Poisson
- (2) Neg Bin I
- (3) Neg Bin II

## Poisson

$$P_a(y_i = y) = \frac{\exp\{-\lambda_i\} \lambda_i^y}{y!} \quad y = 0, 1, 2, 3, \dots$$

$$E(y_i) = \lambda_i$$

$$\text{Var}(y_i) = \lambda_i$$

equidispersion

Then, model  $\lambda_i = \exp\{\gamma_i^T \beta\}$

Hence,  $E(y_i | \gamma_i) = \exp\{\gamma_i^T \beta\}$

$$\text{Var}(y_i | \gamma_i) = \exp\{\gamma_i^T \beta\}$$

$$\begin{aligned}
 (2) \quad \ell = \log L &= \sum_{i=1}^n \left[ -\lambda_i + y_i \log(\lambda_i) - \log y_i! \right] \\
 &= \sum_{i=1}^n \left( -\exp\{\gamma_i^T \beta\} + y_i (\gamma_i^T \beta) - \log y_i! \right)
 \end{aligned}$$

For computational reasons, drop  $\log y_i!$   
 (it does not depend on  $\beta$  anyway)

$$\begin{aligned}
 \frac{d\ell}{d\beta} &= \sum_{i=1}^n (y_i \gamma_i - \exp\{\gamma_i^T \beta\} \gamma_i) \\
 &= \sum_{i=1}^n (y_i - \exp\{\gamma_i^T \beta\}) \gamma_i = 0 \quad (3) \\
 &= \sum_{i=1}^n \varepsilon_i \gamma_i = 0 \quad (3.a)
 \end{aligned}$$

$$E(y_i | \gamma_i) = \exp\{\gamma_i^T \beta\}$$

$$\Rightarrow y_i = \exp\{\gamma_i^T \beta\} + \varepsilon_i$$

$$E(\varepsilon_i | \gamma_i) = 0$$

(3) and (3a) are sample  
 Moment conditions. Consequently,  
 what ever maximizes (2)  
 is consistent as long as (1)  
 is true, (even if  $y_i$  is NOT  
 Poisson distributed). An estimator  
 that possesses this quality is  
~~that~~  
 called

Quasi-Maximum likelihood.

An estimator based on the wrong  
 likelihood is consistent if the  
 moment conditions implied by the  
 likelihood eq's (f.o.c.) are satisfied.

The Trick is to estimate the correct covariance matrix (sandwich)

$$\text{Cov}(\hat{\beta}_{MLE}) = \mathbf{I}(\beta)^{-1} = \left[ E(\exp\{\gamma_i^T \beta\} \gamma_i \gamma_i^T) \right]^{-1}$$

$$\text{Cov}(\hat{\beta}_{QMLE}) = \mathbf{I}(\beta)^{-1} \mathbf{J}(\beta) \mathbf{I}(\beta)^{-1}$$

$$\begin{aligned} \mathbf{J}(\beta) &= E[(y_i - \exp\{\gamma_i^T \beta\})^2 \gamma_i \gamma_i^T] \\ &= E[\varepsilon_i^2 \gamma_i \gamma_i^T] \end{aligned}$$

Basically, this looks like White's sandwich we used in linear regression models.

Replace expectations with sample averages and coeff.  $\beta$ , with their MLE (or QMLE).

Notes: Fine for estimating coeffs,  
But not useful for estimating  
Probabilities.

e.g.  $Pr(5 \text{ patents in year})$   
we could estimate the Avg  
# in a year.

Robust will allow overdispersion  
in est of parameters,  $\overline{Pr}$  weight.

$$\text{i.e. } V(y_i | x_i) = E(\varepsilon_i^2 | x_i) > \exp(x_i^T \beta)$$

### Neg Bin I

$$Var(y_i | x_i) = (1 + \delta^2) \exp\{x_i^T \beta\} \quad (4)$$

only consistent if (4) valid. ~~QMLE~~  
NO QMLE.

### Neg Bin II

$$Var(y_i | x_i) = (1 + \alpha^2 \exp\{x_i^T \beta\}) \exp\{x_i^T \beta\}$$

$$\alpha^2 > 0$$

(1) Poisson Model

- equi dispersion

(2) Neg Bin I

- variance exceeds the mean, but ratio same for all obs  $(1 + \delta^2)$

(3) Neg Bin II

- Variance increases as the <sup>cons</sup> mean increases

Marginal Effects

$$\frac{\partial E(y_i | x_i)}{\partial x_{ik}} = \exp\{x_i^T \beta\} \beta_k$$

same sign.

Semi elasticity

$$= \frac{\partial E(y_i | x_i)}{\partial x_{ik}} \cdot \frac{1}{E(y_i | x_i)} = \beta_k$$

Relative change in the conditional mean if the  $k^{\text{th}}$  regressor changes by 1 unit (C.P.)

For Binary Regressor

$$\frac{E\{y_i | x_{ik} = 1, x_i\}}{E\{y_i | x_{ik} = 0, x_i\}} = \exp\{\beta_k\}$$

conditional mean is  $\exp\{\beta_k\}$   
larger if  $x_{ik} = 1$  instead of 0.

See Verbeek