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opened on: 9 Feb 2010, 09:34:23

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. use http://www.stata-press.com/data/r11/nlswork, clear
(National Longitudinal Survey. Young Women 14-26 years of age in 1968)

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```

. xtreg ln_w grade age c.age#c.age ttl_exp c.ttl_exp#c.ttl_exp tenure c.tenure#
> c.tenure 2.race not_smsa south, be

```

```

Between regression (regression on group means) Number of obs = 28091
Group variable: idcode Number of groups = 4697

```

```

R-sq: within = 0.1591 Obs per group: min = 1
      between = 0.4900 avg = 6.0
      overall = 0.3695 max = 15

```

```

sd(u_i + avg(e_i.))= .3036114 F(10,4686) = 450.23
Prob > F = 0.0000

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```

ln_wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
grade	.0607602	.0020006	30.37	0.000	.0568382	.0646822
age	.0323158	.0087251	3.70	0.000	.0152105	.0494211
c.age#c.age	-.0005997	.0001429	-4.20	0.000	-.0008799	-.0003194
ttl_exp	.0138853	.0056749	2.45	0.014	.0027598	.0250108
c.ttl_exp#						
c.ttl_exp	.0007342	.0003267	2.25	0.025	.0000936	.0013747
tenure	.0698419	.0060729	11.50	0.000	.0579361	.0817476
c.tenure#						
c.tenure	-.0028756	.0004098	-7.02	0.000	-.0036789	-.0020722
2.race	-.0564167	.0105131	-5.37	0.000	-.0770272	-.0358061
not_smsa	-.1860406	.0112495	-16.54	0.000	-.2080949	-.1639862
south	-.0993378	.010136	-9.80	0.000	-.1192091	-.0794665
_cons	.3339113	.1210434	2.76	0.006	.0966093	.5712133

```

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```

```
. xtreg ln_w grade age c.age#c.age ttl_exp c.ttl_exp#c.ttl_exp tenure c.tenure#
> c.tenure 2.race not_smsa south, fe
note: grade omitted because of collinearity
note: 2.race omitted because of collinearity
```

```
Fixed-effects (within) regression      Number of obs      =      28091
Group variable: idcode                 Number of groups   =       4697

R-sq:  within = 0.1727                 Obs per group: min =         1
      between = 0.3505                   avg =                6.0
      overall = 0.2625                   max =                15

corr(u_i, Xb) = 0.1936                 F(8,23386)        =      610.12
                                           Prob > F          =      0.0000
```

ln_wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
grade	(omitted)					
age	.0359987	.0033864	10.63	0.000	.0293611	.0426362
c.age#c.age	-.000723	.0000533	-13.58	0.000	-.0008274	-.0006186
ttl_exp	.0334668	.0029653	11.29	0.000	.0276545	.039279
c.ttl_exp#						
c.ttl_exp	.0002163	.0001277	1.69	0.090	-.0000341	.0004666
tenure	.0357539	.0018487	19.34	0.000	.0321303	.0393775
c.tenure#						
c.tenure	-.0019701	.000125	-15.76	0.000	-.0022151	-.0017251
2.race	(omitted)					
not_smsa	-.0890108	.0095316	-9.34	0.000	-.1076933	-.0703282
south	-.0606309	.0109319	-5.55	0.000	-.0820582	-.0392036
_cons	1.03732	.0485546	21.36	0.000	.9421496	1.13249
sigma_u	.35562203					
sigma_e	.29068923					
rho	.59946283	(fraction of variance due to u_i)				

```
F test that all u_i=0:      F(4696, 23386) =      5.13      Prob > F = 0.0000
```

```
. xtreg ln_w grade age c.age#c.age ttl_exp c.ttl_exp#c.ttl_exp tenure c.tenure#
> c.tenure 2.race not_smsa south, fe vce(robust)
note: grade omitted because of collinearity
note: 2.race omitted because of collinearity
```

```
Fixed-effects (within) regression      Number of obs      =      28091
Group variable: idcode                 Number of groups   =       4697

R-sq:  within = 0.1727                 Obs per group: min =         1
      between = 0.3505                    avg =         6.0
      overall = 0.2625                    max =        15

corr(u_i, Xb) = 0.1936                  F(8,4696)          =      273.86
                                          Prob > F           =      0.0000
```

(Std. Err. adjusted for 4697 clusters in idcode)

ln_wage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
grade	(omitted)					
age	.0359987	.0052407	6.87	0.000	.0257243	.046273
c.age#c.age	-.000723	.0000845	-8.56	0.000	-.0008887	-.0005573
ttl_exp	.0334668	.004069	8.22	0.000	.0254896	.0414439
c.ttl_exp#						
c.ttl_exp	.0002163	.0001763	1.23	0.220	-.0001294	.0005619
tenure	.0357539	.0024683	14.49	0.000	.0309148	.040593
c.tenure#						
c.tenure	-.0019701	.0001696	-11.62	0.000	-.0023026	-.0016376
2.race	(omitted)					
not_smsa	-.0890108	.0137629	-6.47	0.000	-.1159926	-.062029
south	-.0606309	.0163366	-3.71	0.000	-.0926583	-.0286035
_cons	1.03732	.0739644	14.02	0.000	.8923149	1.182325
sigma_u	.35562203					
sigma_e	.29068923					
rho	.59946283	(fraction of variance due to u_i)				

```
. xtreg ln_w grade age c.age#c.age ttl_exp c.ttl_exp#c.ttl_exp tenure c.tenure#
> c.tenure 2.race not_smsa south, re theta
```

```
Random-effects GLS regression           Number of obs       =       28091
Group variable: idcode                  Number of groups     =        4697
```

```
R-sq:  within = 0.1715                   Obs per group: min =         1
      between = 0.4784                               avg =         6.0
      overall = 0.3708                               max =        15
```

```
Random effects u_i ~ Gaussian           Wald chi2(10)       =       9244.74
corr(u_i, X) = 0 (assumed)              Prob > chi2         =        0.0000
```

```
----- theta -----
min      5%      median      95%      max
0.2520   0.2520   0.5499   0.7016   0.7206
```

ln_wage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
grade	.0646499	.0017812	36.30	0.000	.0611589	.0681409
age	.0368059	.0031195	11.80	0.000	.0306918	.0429201
c.age#c.age	-.0007133	.000005	-14.27	0.000	-.0008113	-.0006153
ttl_exp	.0290208	.002422	11.98	0.000	.0242739	.0337678
c.ttl_exp# c.ttl_exp	.0003049	.0001162	2.62	0.009	.000077	.0005327
tenure	.0392519	.0017554	22.36	0.000	.0358113	.0426925
c.tenure# c.tenure	-.0020035	.0001193	-16.80	0.000	-.0022373	-.0017697
2.race	-.053053	.0099926	-5.31	0.000	-.0726381	-.0334679
not_smsa	-.1308252	.0071751	-18.23	0.000	-.1448881	-.1167622
south	-.0868922	.0073032	-11.90	0.000	-.1012062	-.0725781
_cons	.2387207	.049469	4.83	0.000	.1417633	.3356781
sigma_u	.25790526					
sigma_e	.29068923					
rho	.44045273	(fraction of variance due to u_i)				

```
. xtreg ln_w grade age c.age#c.age ttl_exp c.ttl_exp#c.ttl_exp tenure c.tenure#
> c.tenure 2.race not_smsa south, mle
```

Fitting constant-only model:

```
Iteration 0: log likelihood = -13690.161
Iteration 1: log likelihood = -12819.317
Iteration 2: log likelihood = -12662.039
Iteration 3: log likelihood = -12649.744
Iteration 4: log likelihood = -12649.614
Iteration 5: log likelihood = -12649.614
```

Fitting full model:

```
Iteration 0: log likelihood = -8922.145
Iteration 1: log likelihood = -8853.6409
Iteration 2: log likelihood = -8853.4255
Iteration 3: log likelihood = -8853.4254
```

```
Random-effects ML regression      Number of obs      =      28091
Group variable: idcode           Number of groups   =      4697
```

```
Random effects u_i ~ Gaussian    Obs per group: min =      1
                                   avg =      6.0
                                   max =      15
```

```
LR chi2(10) = 7592.38
Prob > chi2 = 0.0000
Log likelihood = -8853.4254
```

ln_wage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
grade	.0646093	.0017372	37.19	0.000	.0612044 .0680142
age	.0368531	.0031226	11.80	0.000	.030733 .0429732
c.age#c.age	-.0007132	.0000501	-14.24	0.000	-.0008113 -.000615
ttl_exp	.0288196	.0024143	11.94	0.000	.0240877 .0335515
c.ttl_exp#					
c.ttl_exp	.000309	.0001163	2.66	0.008	.0000811 .0005369
tenure	.0394371	.0017604	22.40	0.000	.0359868 .0428875
c.tenure#					
c.tenure	-.0020052	.0001195	-16.77	0.000	-.0022395 -.0017709
2.race	-.0533394	.0097338	-5.48	0.000	-.0724172 -.0342615
not_smsa	-.1323433	.0071322	-18.56	0.000	-.1463221 -.1183644
south	-.0875599	.0072143	-12.14	0.000	-.1016998 -.0734201
_cons	.2390837	.0491902	4.86	0.000	.1426727 .3354947
/sigma_u	.2485556	.0035017			.2417863 .2555144
/sigma_e	.2918458	.001352			.289208 .2945076
rho	.4204033	.0074828			.4057959 .4351212

Likelihood-ratio test of sigma\_u=0: chibar2(01)= 7339.84 Prob>=chibar2 = 0.000

```
. reg ln_w grade age c.age#c.age ttl_exp c.ttl_exp#c.ttl_exp tenure c.tenure#c.tenure#
> tenure 2.race not_smsa south, cluster(idcode)
```

Linear regression

```
Number of obs = 28091
F( 10, 4696) = 560.52
Prob > F      = 0.0000
R-squared     = 0.3745
Root MSE     = .37797
```

(Std. Err. adjusted for 4697 clusters in idcode)

ln_wage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
grade	.0629238	.002076	30.31	0.000	.0588539	.0669937
age	.038598	.0049782	7.75	0.000	.0288385	.0483575
c.age#c.age	-.0007082	.0000839	-8.44	0.000	-.0008726	-.0005437
ttl_exp	.0211279	.0036516	5.79	0.000	.0139691	.0282867
c.ttl_exp#c.ttl_exp	.0004473	.0002113	2.12	0.034	.0000332	.0008615
tenure	.0473687	.0028918	16.38	0.000	.0416995	.053038
c.tenure#c.tenure	-.002027	.0001984	-10.22	0.000	-.0024161	-.001638
2.race	-.0699386	.0098805	-7.08	0.000	-.089309	-.0505681
not_smsa	-.1720455	.009528	-18.06	0.000	-.1907248	-.1533662
south	-.1003387	.0088154	-11.38	0.000	-.117621	-.0830565
_cons	.2472833	.0704249	3.51	0.000	.1092175	.385349

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. log close
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name: <unnamed>
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log: H:\Documents and Settings\Lee\My Documents\Document\stata\panel.log
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log type: text
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closed on: 9 Feb 2010, 09:38:20
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## Assessing goodness of fit

$R^2$  is a popular measure of goodness of fit in ordinary regression. In our case, given  $\hat{\alpha}$  and  $\hat{\beta}$  estimates of  $\alpha$  and  $\beta$ , we can assess the goodness of fit with respect to (1), (2), or (3). The prediction equations are, respectively,

$$\hat{y}_{it} = \hat{\alpha} + \mathbf{x}_{it}\hat{\beta} \quad (1''')$$

$$\hat{\bar{y}}_i = \hat{\alpha} + \bar{\mathbf{x}}_i\hat{\beta} \quad (2''')$$

$$\hat{\tilde{y}}_{it} = (\hat{y}_{it} - \hat{\bar{y}}_i) = (\mathbf{x}_{it} - \bar{\mathbf{x}}_i)\hat{\beta} \quad (3''')$$

`xtreg` reports “ $R$ -squares” corresponding to these three equations.  $R$ -squares is in quotes because the  $R$ -squares reported do not have all the properties of the OLS  $R^2$ .

The ordinary properties of  $R^2$  include being equal to the squared correlation between  $\hat{y}$  and  $y$  and being equal to the fraction of the variation in  $y$  explained by  $\hat{y}$ —formally defined as  $\text{Var}(\hat{y})/\text{Var}(y)$ . The identity of the definitions is from a special property of the OLS estimates; in general, given a prediction  $\hat{y}$  for  $y$ , the squared correlation is not equal to the ratio of the variances, and the ratio of the variances is not required to be less than 1.

`xtreg` reports  $R^2$  values calculated as correlations squared, calling them  $R^2$  overall, corresponding to (1''');  $R^2$  between, corresponding to (2'''); and  $R^2$  within, corresponding to (3'''). In fact, you can think of each of these three numbers as having all the properties of ordinary  $R^2$ s, if you bear in mind that the prediction being judged is not  $\hat{y}_{it}$ ,  $\hat{\bar{y}}_i$ , and  $\hat{\tilde{y}}_{it}$ , but  $\gamma_1\hat{y}_{it}$  from the regression  $y_{it} = \gamma_1\hat{y}_{it}$ ;  $\gamma_2\hat{\bar{y}}_i$  from the regression  $\bar{y}_i = \gamma_2\hat{\bar{y}}_i$ ; and  $\gamma_3\hat{\tilde{y}}_{it}$  from  $\tilde{y}_{it} = \gamma_3\hat{\tilde{y}}_{it}$ .

In particular, `xtreg, be` obtains its estimates by performing OLS on (2), and therefore its reported  $R^2$  between is an ordinary  $R^2$ . The other two reported  $R^2$ s are merely correlations squared, or, if you prefer,  $R^2$ s from the second-round regressions  $y_{it} = \gamma_{11}\hat{y}_{it}$  and  $\tilde{y}_{it} = \gamma_{13}\hat{\tilde{y}}_{it}$ .

`xtreg, fe` obtains its estimates by performing OLS on (3), so its reported  $R^2$  within is an ordinary  $R^2$ . As with `be`, the other  $R^2$ s are correlations squared, or, if you prefer,  $R^2$ s from the second-round regressions  $\bar{y}_i = \gamma_{22}\hat{\bar{y}}_i$  and, as with `be`,  $\tilde{y}_{it} = \gamma_{23}\hat{\tilde{y}}_{it}$ .

`xtreg, re` obtains its estimates by performing OLS on (4); none of the  $R^2$ s corresponding to (1'''), (2'''), or (3''') correspond directly to this estimator (the “relevant”  $R^2$  is the one corresponding to (4)). All three reported  $R^2$ s are correlations squared, or, if you prefer, from second-round regressions.