

## IIA

A disadvantage of MNL and LL is that the odds ratios are independent of the other alternatives

TAKE 3 ALTERNATIVES :

~~WAYS~~ Commute =  $\left\{ \begin{array}{l} \text{Alone} \quad (A) \\ \text{BUS} \quad (B) \\ \text{Carpool} \quad (C) \end{array} \right.$

$$\text{Prob}(A) = \frac{e^{-\tau_{PA}}}{1 + e^{-\tau_{PA}} + e^{-\tau_{PB}}}$$

$$\text{Prob}(B) = \frac{e^{-\tau_{PB}}}{1 + e^{-\tau_{PA}} + e^{-\tau_{PB}}}$$

$$\text{Prob}(C) = \frac{1}{1 + e^{-\tau_{PA}} + e^{-\tau_{PB}}}$$

Hence,  $\frac{\text{Prob}(A)}{\text{Prob}(C)} = e^{-\tau_{PA}}$        $\frac{\text{Prob}(B)}{\text{Prob}(C)} = e^{-\tau_{PB}}$

Odds ratios do not depend on other alternatives. ~~Prob~~ Relative Prob of commuting alone is not affected by BUS (i.e.  $\tau_{PB}$ )

Let's say

$$P(A) = 1/2$$

$$P(B) = 1/3$$

$$P(C) = 1/6$$

Odds Ratios.

$$\frac{P(A)}{P(C)} = 3$$

$$\frac{P(B)}{P(C)} = 2$$

Now, suppose we add an irrelevant alternative 4  
alternative.

Red Bus - RB

Blue Bus - BB

Assuming people are in different  
between the Blue & Red colors  
of bus, we'd expect this to have  
no effect on the prob of going alone  
or of carpooling. -

unfortunately, though, in the  
MNL and CL it does.

- How so? Well, the estimate of  $\beta_B$   
depends on the characteristics of the  
individual (MNL) or attributes (relative <sup>which are</sup>)  
They should not be affected by the color  
of the bus.

$$\text{So: } \beta_{RB} = \beta_{BB} = \beta_B$$

$$P_{\text{ROB}}(A) = \frac{e^{\gamma \beta_A}}{1 + e^{\gamma \beta_{RB}} + e^{\gamma \beta_{BB}} + e^{\gamma \beta_C}}$$

$$\begin{aligned} \beta_{RB} &= \beta_{BB} = \beta_A \\ &= \frac{e^{\gamma \beta_A}}{1 + 2e^{\gamma \beta_B} + e^{\gamma \beta_A}} \end{aligned}$$

which is different than before.

also, the odds ratios don't change

$$\frac{P_{\text{ROB}}(A)}{P_{\text{ROB}}(C)} = e^{\gamma \beta_A}$$

$$\frac{P_{\text{ROB}}(BB)}{P_{\text{ROB}}(C)} = \frac{P_{\text{ROB}}(RB)}{P_{\text{ROB}}(C)} = e^{\gamma \beta_B}$$

MNL

$$\begin{aligned} P(A) &= 3/8 \\ P(RB) &= 1/4 \\ P(BB) &= 1/4 \\ P(C) &= 1/8 \end{aligned}$$

Common Sense

$$\begin{aligned} P(A) &= 1/2 \\ P(RB) &= P(BB) = 1/6 \\ P(C) &= 1/6 \end{aligned}$$

When IIA fails, An alternative estimator is needed.

(1) Multivariate Probit

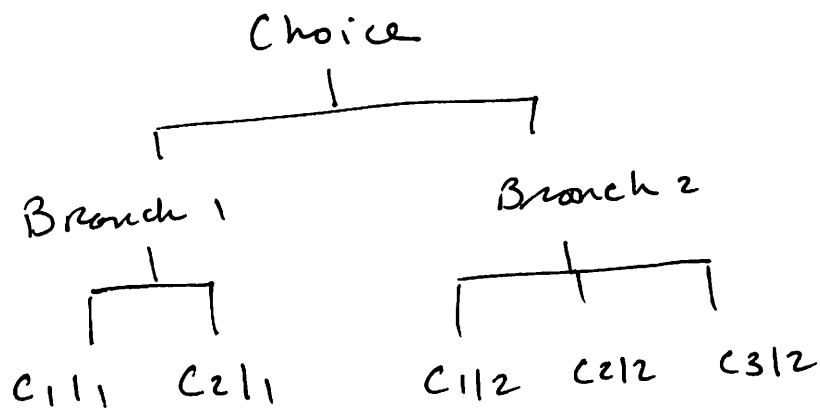
$$U_{ij} = \gamma_{ij}^T \beta + \varepsilon_{ij}$$

$$j = 1, 2, \dots, J$$

$$\varepsilon_{i1}, \varepsilon_{i2}, \dots, \varepsilon_{ij} \sim N(\mathbf{0}, \Sigma)$$

In small dimensional problems  
( $J$  small), this is computationally  
feasible.

(2) Nested logit



"twigs"

## TEST for IIA

Hausman & McFadden (1984)

If an alternative really is irrelevant, then omitting it will not change the parameter estimates very much, if any. Excluding an irrelevant choice won't be efficient & But it is consistent.

If the odds ratios do depend on the alternatives, then the parameter estimates obtained when these choices are excluded are inconsistent.

Let  $\hat{\beta}_s$ ,  $\hat{V}_s$  be the estimates of  $\beta_s$  with subset of the choices, and est. covariance

Let  $\hat{\beta}_f$ ,  $\hat{V}_f$  be the estimates

of  $\beta$  and  $\text{Cov}(\hat{\beta}_f)$  using the full

set

$$\chi^2_k = \left( \hat{\beta}_s - \hat{\beta}_f \right)^T \left[ \hat{V}_s - \hat{V}_f \right]^{-1} \left( \hat{\beta}_s - \hat{\beta}_f \right)$$

$\begin{matrix} 1 \times k & & k \times k & & k \times 1 \end{matrix}$

$\hat{\beta}_s, \hat{\beta}_f$  are  $k \times 1$

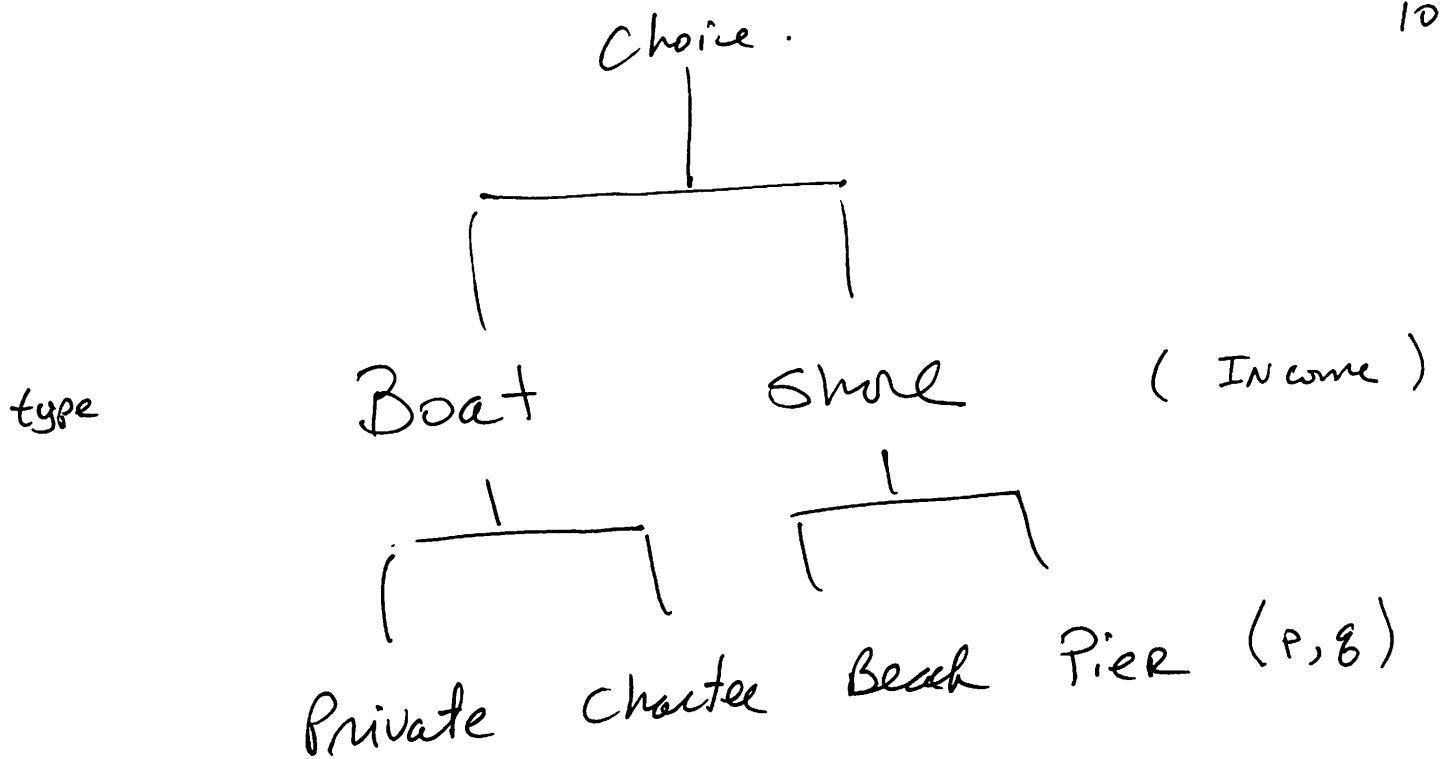
$\chi^2_k$  if  $H_0$ : True.

$H_0$ : <sup>one (a more)</sup> choices irrelevant

$H_A$ : not irrelevant.

$$\ln L = \sum_{i=1}^n \ln [\text{Prob}(\text{twig} | \text{branch})_i \times \text{Prob}(\text{branch})_i]$$

Recall  $\text{Prob}(A, B) = \text{Prob}(A|B) \text{Prob}(B)$



IIA applies within each Branch  
But not across Branches.

nlogitgen type = fishmode (shore: pier | beach,  
boat: private | charter)

nlogittree fishmode type, choice(d)

```

. * estimate the model
. nlogit d p q || type:, base(shore) || fishmode: income, case(id) notree nolog

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RUM-consistent nested logit regression      Number of obs      =      4728
Case variable: id                          Number of cases     =      1182

Alternative variable: fishmode              Alts per case: min =      4
                                              avg   =      4.0
                                              max   =      4

Log likelihood = -1192.4236                 wald chi2(5)       =      212.37
                                              Prob > chi2        =      0.0000

```

	d	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
fishmode	p	-.0267625	.0018937	-14.13	0.000	-.0304741	-.023051
	q	1.340078	.3080391	4.35	0.000	.7363329	1.943824

## fishmode equations

beach	income	(base)					
_cons		(base)					
charter	income	-8.402969	78.34197	-0.11	0.915	-161.9504	145.1445
	_cons	69.96747	558.6645	0.13	0.900	-1024.995	1164.93
pier	income	-9.458769	80.28494	-0.12	0.906	-166.8144	147.8968
	_cons	58.94565	500.589	0.12	0.906	-922.1908	1040.082
private	income	-1.634704	8.583156	-0.19	0.849	-18.45738	15.18797
	_cons	37.51814	230.7252	0.16	0.871	-414.695	489.7313

## dissimilarity parameters

type	/shore_tau	718.2333			1491.178
/boat_tau	52.56573	542.7797			1116.394

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LR test for IIA (tau = 1):                chi2(2) =      45.43   Prob > chi2 = 0.0000

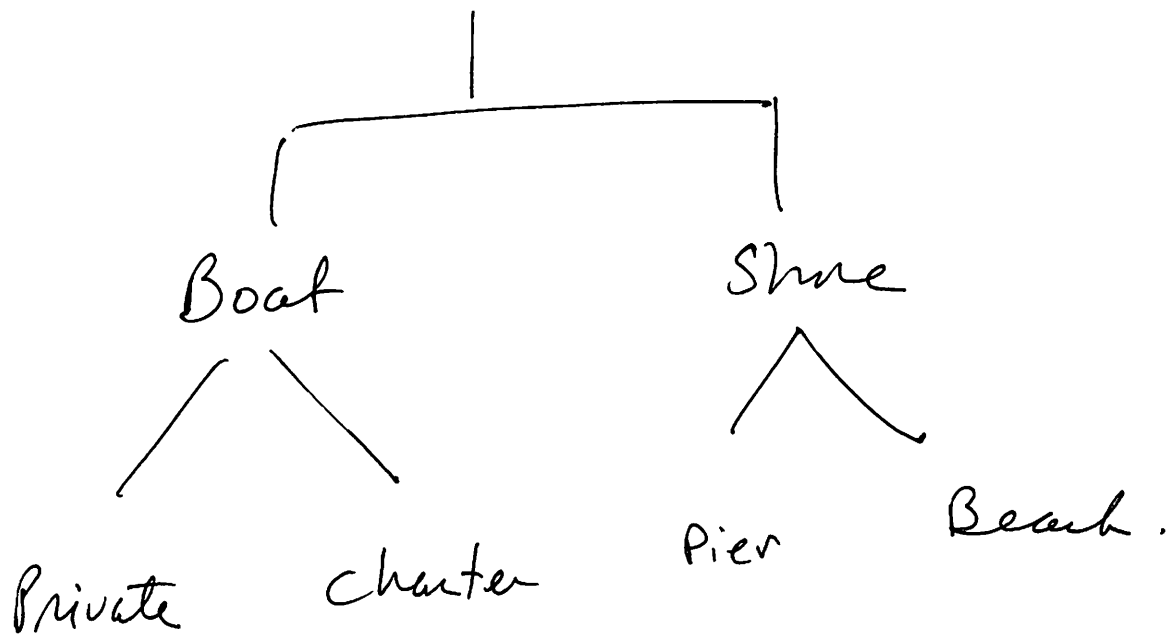
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. predict plevel1 plevel2, pr
. tabulate fishmode, summarize(plevel2)

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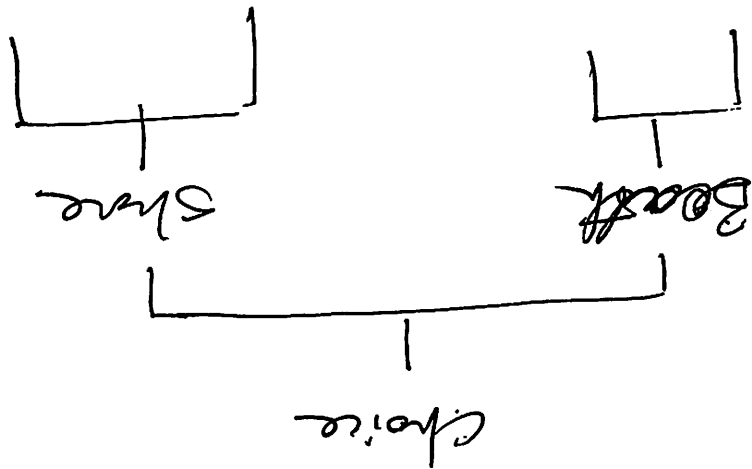
fishmode	Summary of Pr(fishmode alternatives)		
	Mean	Std. Dev.	Freq.
beach	.1132351	.13335995	1182
charter	.38070838	.15724242	1182
pier	.15072736	.16982071	1182
private	.35532917	.16444521	1182
Total	.25	.19690069	4728



in nested Model, The errors within each Branch are correlated.

$\epsilon_{ij}$ : Private & charter       $\epsilon_{ij}$ : Pier & Beach.

CL is a special case where all errors are independent.



```

1 use "C:\DATA\stata\mus\mus15data.dta", clear
2
3 log using "H:\Documents and Settings\Lee\My Documents\Document\stata\asclogit.smcl",
  replace
4
5 generate id = _n
6 list mode price pbeach ppier pprivate pcharter in 1/4, clean
7
8 * Multinomial Logit -- case specific vars only in short form
9 mlogit mode income, baseoutcome(1)
10
11 * Convert data to long form for asclogit
12 reshape long d p q, i(id) j(fishmode beach pier private charter) string
13 list in 1/4
14
15 * Drop the variables that don't make sense in long form
16 drop mode price crate
17
18 * Conditional logit with case and alternative specific variables
19 asclogit d p q, case(id) alternatives(fishmode) casevars(income)
20
21 * Predicted probabilities of choice of each mode and compare to actual freqs
22 predict pasclogit, pr
23 table fishmode, contents(mean d mean pasclogit sd pasclogit) cellwidth(15)
24
25 * Marginal effect at mean of change in price
26 estat mfx, varlist(p)
27
28 * MNL is CL with no alternative specific regressors
29 asclogit d, case(id) alternatives(fishmode) casevars(income) basealternative(beach)
30
31 * Nested Logit
32 * define the tree
33 nlogitgen type = fishmode(shore: pier | beach, boat: private | charter)
34
35 * check the tree
36 nlogittree fishmode type, choice(d)
37
38 * estimate the model
39 nlogit d p q || type:, base(shore) || fishmode: income, case(id) notree nolog
40
41 predict plevel1 plevel2, pr
42 tabulate fishmode, summarize(plevel2)

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