

Other Binary Choice Models

Logit

$$\Pr(y=1) = \frac{\exp\{\gamma_i^T \beta\}}{1 + \exp\{\gamma_i^T \beta\}} \equiv \Lambda(\gamma_i^T \beta)$$

Λ is the logistic CDF

The logit p.d.f.

$$\begin{aligned} \frac{d \Lambda(\gamma_i^T \beta)}{d(\gamma_i^T \beta)} &= \frac{\exp\{\gamma_i^T \beta\}}{[1 + \exp(\gamma_i^T \beta)]^2} \\ &= \Lambda(\gamma_i^T \beta) [1 - \Lambda(\gamma_i^T \beta)] \end{aligned}$$

\therefore The marginal effects

$$\frac{\partial E(y|\gamma)}{\partial \gamma} = \Lambda(\gamma^T \beta) (1 - \Lambda(\gamma^T \beta)) \beta$$

Hessian:

$$\frac{\partial^2 \ln L}{\partial \beta \partial \beta^T} = - \sum \Lambda_i (1 - \Lambda_i) \gamma_i \gamma_i^T$$

$\Lambda_i \equiv \Lambda(\gamma_i^T \beta)$

$\frac{N}{n} t_1$

The tails of a logit dist are heavier than those of the normal.

If the index $\eta_i^T \beta$ $i=1, \dots, n$

lie between ± 1.2 then the

results will be very similar.

The biggest differences occur when

either $\eta_i^T \beta$ is very large or small.

Predictions will differ especially

if (1) there are very few 0's or 1's and (2) if there is very wide variation in $\eta_i^T \beta$.

Logit & Probit are Both symmetric.

Non symmetric

GUMBEL

$$\text{Prob}(y=1 | \underline{x}) = \exp\{-\exp(-\underline{x}^T \beta)\}$$

Complementary log-log

$$\text{Prob}(y=1 | \underline{x}) = 1 - \exp\{-\exp\{\underline{x}^T \beta\}\}$$

Skewed logit

$$\text{Prob}(y=1) = \frac{1}{[1 + \exp\{\underline{x}^T \beta\}]^\alpha}$$

STATA: The main reason to use this model is that the maximum prob effect of the regressors on the prob of success is not constrained to be when $p = .5$. Instead, it's used when that max occurs ~~at~~ between .3 and .6

- * How you code 1 & 0 matters
- * you could use "Power logit" if $p > .63$
- * you would reverse the signs and talk about the marginal effect of \underline{x} on failure

Model with interactions

$$\text{Prob}(\text{DocVis} = 0) = \Lambda(\beta_1 + \beta_2 \text{Age} + \beta_3 \text{Inc.} \\ + \beta_4 \text{Kids} + \beta_5 \text{Edec} + \beta_6 \text{MAER} + \beta_7 \text{Age} * \text{Edec})$$

$$\frac{\partial \text{Prob}(\text{DocVis} = 0)}{\partial \text{Age}} = \Lambda(\gamma\beta) (1 - \Lambda(\gamma\beta)) (\beta_2 + \beta_7 \text{Edec})$$

$$\frac{\partial \text{Prob}(\text{DocVis} = 0)}{\partial \text{Edec}} = \Lambda(\gamma\beta) [1 - \Lambda(\gamma\beta)] (\beta_5 + \beta_7 \text{Age})$$

So far so good

But,

$$\frac{\partial^2 \text{Prob}(\text{DocVis})}{\partial \text{Age} \partial \text{Edec}} = \Lambda(\gamma\beta) [1 - \Lambda(\gamma\beta)] \beta_7 + \\ \Lambda(1 - \Lambda) (\beta_5 + \beta_7 A) (\beta_2 + \beta_7 \epsilon) \\ - 2\Lambda(\Lambda(1 - \Lambda)) (\beta_5 + \beta_7 A) \\ \times (\beta_2 + \beta_7 \epsilon) \\ = \Lambda(1 - \Lambda) \beta_7 + \Lambda(1 - \Lambda) [1 - 2\Lambda] \\ \times (\beta_2 + \beta_7 \text{Edec}) (\beta_5 + \beta_7 \text{Age})$$

Fit

$$\text{pseudo-} R^2 = 1 - \frac{\ln L}{\ln L_0}$$

$\ln L$ is log likelihood on net

$\ln L_0$ is log likelihood with ACC
"slopes" set to 0, i.e.,

$$\beta_2 = \beta_3 = \dots = \beta_k = 0$$

This stat = 0 if $\beta_2 = \dots = \beta_k = 0$

But can never quite = 1 (unless

There is something seriously wrong
with your model i.e., $\gamma \beta \rightarrow \infty$

Another

$$R_{bc}^2 = \frac{1}{n} \sum_{i=1}^n y_i \hat{F}_i + (1 - y_i)(1 - \hat{F}_i)$$

Avg prob of correct prediction

However, in badly balanced
samples the less freq outcome
is seldom predicted.

Cramer.

$$\lambda = (\text{Average } \hat{F} | y_i = 1) - \text{Avg}$$

$$(\text{Average } \hat{F} | y_i = 0)$$

This heavily penalizes incorrect predictions.

TABLE

Might want to do Δ \cdot q $p = .5$
rule.