

Fixed Effects Model

$$\underline{y} = X_s \underline{\beta} + D \underline{\alpha} + \underline{\varepsilon}$$

$$D = I_n \otimes \underline{1}_T$$

FE estimator:

use M_D to remove dummy vars from model.

$$M_D = I - P_D \quad P_D = D(D^T D)^{-1} D^T$$

$$\begin{aligned} M_D \underline{y} &= M_D X_s \underline{\beta} + M_D D \underline{\alpha} + M_D \underline{\varepsilon} \\ &= M_D X_s \underline{\beta} + M_D \underline{\varepsilon} \end{aligned}$$

$$\hat{\underline{\beta}}_{FE} = (X_s^T M_D X_s)^{-1} X_s^T M_D \underline{y}$$

Between Effects Model

$$\bar{y}_{i.} = \bar{x}_{i.}^T \beta + \alpha + \bar{e}_{i.} \quad i=1, \dots, n$$

This Model can be obtained
using $P_D = D(D^T D)^{-1} D^T$

Recall:

$$\begin{aligned} P_D &= (I_n \otimes \bar{i}_{.}) \left[(I_n \otimes \bar{i}_{.})^T (I_n \otimes \bar{i}_{.}) \right]^{-1} (I_n \otimes \bar{i}_{.})^T \\ &= I_n \otimes \frac{\bar{i}_{.} \bar{i}_{.}^T}{T} \end{aligned}$$

Model

$$P_D \underline{y} = P_D X \underline{\beta} + \underbrace{P_D D}_{\alpha} + P_D \underline{\epsilon}$$

$$\underline{P}_D D = D$$

But α are not separately identified

NOTES:

Asymptotics

n - # individuals or groups

T - time periods

nT - Total sample size
(balanced)

1) Suppose n is fixed. As $T \rightarrow \infty$

The number ~~of~~ of fixed effects

Remains constant. So, we

have more obs to estimate

both slopes and α_i as $T \rightarrow \infty$

\Rightarrow Standard Asympt Theory
Applies

2) Suppose \underline{n} is $\underline{\underline{\rightarrow \infty}}$ with \underline{T} fixed

\Rightarrow # parameters to estimate
is also converging to ∞

Then $\hat{\alpha}_i$ are NOT
consistent in fixed effects
Estimate.

3) Even if $n \rightarrow \infty$ and T -fixed

$\hat{\beta}_{FE}$ is consistent for β

(miraculous!)!

4) An estimator of σ_u^2 is
 The variance in \hat{d}_i
 Sample

Obviously, the larger the
 number of \hat{d}_i the better (finite
 (except if $n \rightarrow \infty$.)

Cluster Std. Errors

Consider The Pooled Model

$$y_{it} = \gamma_{it}' \beta + \alpha + \varepsilon_{it}$$

if $\alpha_i \neq \alpha$ for all i

Then LS is inconsistent.

for fixed effects Models.

If α_i correlated with x_s :

If α_i are random, LS might be consistent. Basically there is autocorrelation in the

model
$$\text{Cov}(\alpha_{it}, \alpha_{is}) = \sigma_\alpha^2 \quad \begin{matrix} t \neq s \\ i = j \end{matrix}$$

AND The usual OLS Result Applies.

Look at the i^{th} Group α

~~the~~ individual

$$y_i = \cancel{X_i \beta} + \mu_i$$

$$\mu_i = d_i + \epsilon_i$$

β includes the constant

X_i ~~matrix~~ matrix and include i_T

$$X_i = \begin{bmatrix} i_T & \vdots & X_{si} \end{bmatrix}$$

$$\text{Cov}(\mu_i) = \Sigma = \begin{bmatrix} \sigma_\epsilon^2 + \sigma_\alpha^2 & \sigma_\alpha^2 & \dots & \sigma_\alpha^2 \\ \sigma_\alpha^2 & \dots & \dots & \sigma_\alpha^2 \\ \vdots & \dots & \dots & \sigma_\alpha^2 \\ \sigma_\alpha^2 & \dots & \sigma_\alpha^2 & \sigma_\epsilon^2 + \sigma_\alpha^2 \end{bmatrix}$$

So, there is intragroup autocorrelation

1) Estimate model using LS
(looks like pooled model)

Get LS estimates of residuals
for each group

$$\hat{\tilde{\eta}}_i$$

Est "Cochran" Robert Cov
matrix is

$$\frac{1}{n} \left[\frac{1}{n} \sum_{i=1}^n X_i^T X_i \right]^{-1} \left[\frac{1}{n} \sum_{i=1}^n X_i^T \hat{\tilde{\eta}}_i \hat{\tilde{\eta}}_i^T X_i \right]$$

$$\left[\frac{1}{n} \sum_{i=1}^n X_i^T X_i \right]^{-1}$$

SAND with Cov

are these with

Pooled LS estimates.

$$\underline{i}_{mT} = \underline{i}_m \otimes \underline{i}_T$$

$$P_b = I_m \otimes \underline{i}_T \underline{i}_T^T / T$$

∴

$$P_b \underline{i}_{mT} = \left(I_m \otimes \frac{\underline{i}_T \underline{i}_T^T}{T} \right) (\underline{i}_m \otimes \underline{i}_T)$$

$$= \underline{i}_m \otimes \frac{\underline{i}_T \underline{i}_T^T \underline{i}_T}{T}$$

$$= \underline{i}_m \otimes \underline{i}_T T / T = \underline{i}_m \otimes \underline{i}_T$$

$$= \underline{i}_{mT} !$$

$$y = X\beta + u \quad u \sim (0, \Sigma)$$

use LS with a "Robust"
covariance estimator.

$$\Sigma = \begin{bmatrix} \sigma_1^2 & & & 0 \\ & \sigma_2^2 & & \\ 0 & & \dots & \\ & & & \sigma_T^2 \end{bmatrix}$$

Heteroscedastic.

(1) use LS with
White's cov estimator

(2) use GLS (or FGLS)
with FGLS cov.

$$y_{it} = \gamma_{it} \beta + \alpha + \epsilon_{it}$$

(1) use LS (Pooled)
and Cluster Std errors

(2) GLS or FGLS is the
Random effects estimator.
That specifically models
the intragroup autocorrelations
 σ_α^2 .

Are Effects Random?

$$\alpha_i \sim (0, \sigma_\alpha^2) \Rightarrow \alpha_i \text{ Random}$$

If $\sigma_\alpha^2 = 0$ Then They are NOT
Random.

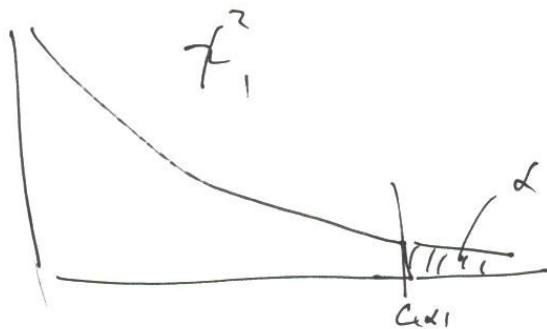
$$H_0: \sigma_\alpha^2 = 0$$

$$H_A: \sigma_\alpha^2 > 0$$

$$\lambda = \frac{NT}{2(T-1)} \left[\frac{\bar{e}^T (I_N \otimes i_T i_T^T) \bar{e}}{\bar{e}^T \bar{e}} - 1 \right]^2 \xrightarrow{d} F_1^2$$

if H_0 True.

\bar{e} residuals from Pooled Regression



Reject if $\lambda > c_{\alpha, 1}$ (Effects Random)

Prediction (RE)

$$(1) E(y_{it} | X_{it})$$

$$\hat{\alpha} + \sum_{k=2}^K \hat{\beta}_k \gamma_{kit}$$

Average over all time periods

? Groups given certain

characteristics (no ind. effects)

(2) Predict y for i^{th} individual | γ_{it}

$$\hat{\alpha} + \sum_{k=2}^K \hat{\beta}_k \gamma_{kit} + \hat{\alpha}_i$$

$$\hat{\alpha}_i = \frac{\hat{\sigma}_\alpha^2}{\hat{\sigma}_i^2} \hat{c}_i^T (y_i - \gamma_i^T \hat{\beta}_{RE})$$

$$= \frac{\hat{\sigma}_\alpha^2}{\hat{\sigma}_i^2 / T} \hat{c}_i^T \frac{(y_i - \gamma_i^T \hat{\beta}_{RE})}{T}$$

$$= \frac{\hat{\sigma}_\alpha^2}{\hat{\sigma}_i^2 / T} \hat{e}_{i0}$$

usage residual in
iD Group

weighted By Proport
of variance explained
By individual, Relat
o Variance of Beta
estimator.
